

Coordination and Self-Organization in Minority Games: Experimental Evidence*

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Abstract

This work presents experimental results on a coordination game in which agents must repeatedly choose between two sides, and a positive fixed payoff is assigned only to agents who pick the minoritarian side. We conduct laboratory experiments in which stationary groups of five players play the game for 100 periods, and manipulate two treatment variables: the amount of ‘memory’ M that players have regarding the game history (i.e., the length of the string of past outcomes that players can see on the screen while choosing), and the amount of information about other players’ past choices. Our results show that, at the aggregate level a quite remarkable degree of coordination is achieved. Moreover providing players with full information about other players’ choice distribution does not appear to improve efficiency significantly. At the individual level, a substantial portion of subjects exhibit ‘inertial’ behavior.

1 Introduction

Many social interaction phenomena imply some advantage that arises from acting in opposition to the modal behavior in a population. For example, if deciding when to leave for vacation with our car, we will be better off avoiding rush hours and travelling when only few other people are on the road. If we want to spend some hours surfing on the web, we will prefer to do that when few other people are surfing so that the connection is faster. If a long-awaited new movie hit is being shown at our local theater, we will prefer to go on those nights in which spending two hours in line for tickets is *least* likely, and so on.

Thomas Schelling in *Micromotives and Macrobehavior* (1978) described and modelled a vast array of social situations that fall under this broad category, from micro-phenomena like the ones just mentioned to far more relevant ones such as racial segregation and house migration. Clearly, what matters in this type of interactive settings is the interplay between the individuals’ expectations and iterated expectations (i.e., expectations of expectations), and aggregate behavior in terms of goodness of coordination. The key to the whole problem in fact consists not simply in trying to guess what the majority will do, and do the opposite, but rather to guess what the majority will think that the majority will do, and so on ad infinitum. Intuitively, the more heterogeneous individual expectations are, the ‘smoother’ - and hence

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efficient - the aggregate outcome of interaction is likely to be (e.g., no traffic jams on the highways, no band-width congestion, etc.), because different beliefs will induce differentiated enough behaviors so as to produce a kind of self-organization at the population level. Too similar expectations, on the contrary, may be likely to result in large overall inefficiency. This point had been first made long ago by B. Arthur (1994) in his ‘El Farol Bar’ model of interaction. Similar coordination problems have been extensively investigated within the experimental economics literature, usually going under the heading of ‘market entry games’ (e.g., Meyer et al. 1992; Ochs 1990, 1995 and references therein).

In this work we present experiments on a very simple repeated game in which the payoff to each player is based on a minority rule. In the Minority Game, first introduced by Challet and Zhang (1997, 1998) a fixed group of N players (where N is odd) must privately and independently choose each round between two actions available to them. The players choosing the action that is chosen by the minority of the players earn a fixed, positive payoff, while players who end up on the majority side earn nothing. Hence, the two alternatives are perfectly symmetric and individual payoff is solely based on how players distribute between them¹. Formally, the model is an N -person non-cooperative game with multiple Nash equilibria in pure strategies; the pure strategy Nash equilibria are Pareto-equivalent and asymmetric with respect to players, in the sense that players differentiate from one another on the basis of the side chosen; further, the game has a unique symmetric equilibrium in mixed strategies, requiring that all agents pick the two sides with equal probability.

This work is aimed at answering the following research questions: what is the degree of efficiency of the system when groups of players interact repeatedly in a Minority Game compared to the benchmark value offered by the mixed strategy equilibrium solution? What is the impact - in terms of efficiency - of varying the amount of information that players have regarding the game history and the past actions of the other players? Do players randomize between rounds, as prescribed by the Nash equilibrium solution? Or are they able to detect and exploit predictable patterns in others’ behavior, when they are given more information?

In order to answer these questions, we conducted experiments in which stationary groups of players played a minority game repeatedly and under different information conditions (see section 3 for a description of the experimental design).

Our preliminary results can be roughly summarized as follows: first, efficiency is higher on average than the benchmark value corresponding to the mixed strategy Nash equilibrium in all treatments, suggesting that a quite remarkable degree of coordination is achieved; second, providing players with more information does not appear to improve efficiency significantly. At the individual level, a substantial portion of subjects exhibit ‘inertial’ behavior, i.e., the tendency to replicate their previous round choice with a higher frequency than the one prescribed by randomizing behavior.

The paper is organized as follows: Section 2 describes the game-theoretic framework and discusses previous related work. Section 3 describes the experimental design, and section 4 analyzes the results. Finally, section 5 offers some concluding remarks and directions for future research.

¹For a large collection of papers concerning both the analytical and numerical explorations of the original minority game and various extensions see also <http://www.unifr.ch/econophysics/>.

2 The Minority Game

The minority game is played repeatedly by a stationary set of N players, where N is an odd number. On each period of the stage game, each player must choose privately and independently between two actions or sides which will label 0 and 1. The payoff function, which is the same for all players is given by

$$\pi_i = \begin{cases} 1 & \text{if } k_i \leq (N-1)/2 \\ 0 & \text{otherwise} \end{cases} \quad (1)$$

where $i \in \{0, 1\}$ and k_i is the number of players choosing i . In descriptive terms, players are rewarded with one point each whenever the side they choose is chosen by the minority, while players who end up on the majority side obtain a null payoff.

It is straightforward to see that the game has $N!/(((N-1)/2)!)^2$ Nash equilibria in pure strategies, in which exactly $(N-1)/2$ players choose either one of the two sides. The pure strategy Nash equilibria are Pareto equivalent and they are not strict, as players on the majority side are just indifferent between deviating and sticking to equilibrium play. Besides, they imply a payoff-asymmetry between players belonging to the different sides, which can rule out the possibility that groups may develop a simple form of tacit coordination based on historical precedent, as it may occur in analogous games (e.g., Meyer et al. 1992). The game also has a unique symmetric mixed-strategy Nash equilibrium, in which each player selects the two actions with equal probability. Finally, there are infinite asymmetric mixed strategy equilibria.

From a game-theoretic standpoint, the game is analogous to the market entry games studied experimentally by Ochs (1990), Meyer et al. (1992), Sundali et al. (1995) and Rapoport et al. (1998). In most of these games, a group of N players must decide whether or not to enter a market that has a certain fixed capacity k , where k is typically $\leq N^2$. Some of the parameters of the payoff function may vary so as to render the ‘staying out’ option more or less attractive to players (in some cases players receive a fixed payoff from not entering). The payoff to entrants usually declines linearly in their number. In all such games there is usually a unique symmetric mixed strategy equilibrium and several pure strategy equilibria. The minority game, although belonging to the same class of coordination games, differs from the previously studied market entry games in some crucial aspects: its equilibrium structure is extremely unstable relative to other games, in that all pure strategy equilibria, whereby players ‘lock in’ in the choice of the same side, besides being non strict, and hence subject to being disrupted by minor deviations, imply a payoff-asymmetry between the two resulting parties which strongly increase their instability; for this reason, MG may be the ideal setup to study the emergence of collective efficiency resting on some forms of dynamic coordination between players. Moreover, while in market entry games the value of the market capacity usually changes from period to period thus changing the probability distribution implied by the mixed strategy equilibrium, in the MG this remains fixed throughout the game and it is very simple, as it assigns equal probability to the two available options; this, combined with the large numbers of rounds, the small group size and the strongly competitive nature of the game, should provide a strong incentive to random behavior and should allow for learning, thus making the unique mixed strategy equilibrium a natural ‘benchmark’ against which to confront actual behavior. Deviations from the mixed strategy hypothesis hence may

²In other cases, as in Meyer et al. (1992), there are actually two markets between which players must allocate their entry decisions.

be interpreted as the result of decision rules - other than randomization - which may depend on the game structure and to the history of interaction rather than as deriving from subjects' 'imperfect' ability to generate random series (e.g., Budescu 1987).

3 The experiment: design and implementation

The information treatments

We conducted experiments in which stationary groups of five players played a minority game for 100 rounds, and we applied a 2×3 factorial design in which the two treatment variables were the amount of 'memory' M that players could have regarding the game history (i.e., the number of past rounds for which information was made visible on the subjects' computer screen), and the amount of information that players had on the past choices of the other players.

A total of 120 subjects participated in the experiment. Four groups of five players participated in the single treatments. A group size of five seems large enough as to prevent subjects from developing repeated game strategies aimed at achieving tacit coordination. Indeed, such a form of dynamic coordination is not trivial to achieve in the minority game even with a small group, because it would require a complex rotation in the chosen side between players³.

Table 1 summarizes the experimental design.

As previously stated, in this work we are mainly interested in studying the impact of information on aggregate efficiency and individual behavior.

Several previous simulation studies on the MG (e.g., Challet and Zhang 1997, 1998) have focussed on the role of the information available to players in determining aggregate efficiency. In most of these studies the amount of information given to artificial agents was varied by changing the value of the 'memory' parameter M , defined as the length of the string of past outcomes that agents were able to retain and process when making their choices. Indeed, these works have shown that the amount of agents' memory (and hence the amount of information on the 'history' of interaction) has a significant impact on the system's degree of self-organization when artificial agents are designed to behave according to experience-based strategies.

In line with previous simulation studies, we also chose to focus on the role of information; moreover, in this study we decided to jointly investigate the impact of both information about the series of past outcomes (i.e., the players' available memory capacity) and information (or lack thereof) about the behavior of single players in the population.

As far as the first parameter is concerned, in the case of human agents memory itself cannot be directly controlled for; hence, a rough 'proxy' for it was devised by varying the number of past outcomes for which information was disclosed and remained visible on the subjects' computer screens while playing. We chose three levels of memory: short ($M = 1$), medium ($M = 4$), and long ($M = 16$). In the last case, we assume that the amount of information made available is considerably more than what individuals are able to process, due to well known human working memory constraints (e.g., Miller 1956).

The second treatment variable was the type of information provided to subjects at each round of play. In the *aggregate* information treatment, subjects were only allowed to know which side (0 or 1) was the winning side at each round (i.e., they only knew the aggregate outcome of interaction at every time step), while in the *full* information treatment, subjects could also see the entire distribution of individual choices in their group. Besides, the choices

³This form of tacit cooperation was in fact never observed in our experiments.

	Memory (M)		
	$M = 1$	$M = 4$	$M = 16$
Information			
Aggregate (A)	4	4	4
Full (F)	4	4	4

Table 1: The 3×2 factorial design of the experiments. Numbers in the cells indicate the number of groups that were assigned to each treatment

steps	$M = 1$		$M = 4$		$M = 16$	
	A	F	A	F	A	F
25	1.31	1.05	.99	1.45	1.25	.91
50	1.25	1.29	.83	1.13	1.29	.71
75	1.19	1.31	.93	.85	.89	.97
100	.99	1.03	.91	.77	.89	.61

Table 2: Values of σ averaged over successive 25 time steps and across groups, reported separately for values of M and for the aggregate (A) and full (F) information treatments

of players always appeared in the same order, and this fact was explicitly stated in the instructions, so subjects were able to keep track of every individual history of choices in their group.

Previous experimental studies have shown that providing information about other players' choices usually has a significant impact on individual behavior in repeated games, although such impact tends to be game-specific and not necessarily efficiency-enhancing. Duffy and Hopkins (2001) showed that when players had full information about others' choices, they converged much more quickly to equilibrium play in a repeated market entry game with strict pure strategy equilibria. On the contrary, providing subjects with full information about others' choices lowered efficiency in a large-group coordination game with multiple Pareto-ranked equilibria (Van Huyck, Battalio and Beil 1990). Hence, information about others' behavior, though used by players to modify their choices, not necessarily leads to better collective outcomes compared to the case in which only minimal or no information is available. In the case of a minority game, it would be reasonable to expect that when players have full information (and such condition is common knowledge), they might try to render their behavior less predictable, knowing that other players could exploit patterns in their choices, and might at the same time try to exploit patterns in others' behavior, provided they can detect any of them. Thus, it is plausible to expect that a full information treatment would increase randomizing behaviour.

Implementation

Subjects on each round had to choose between two actions that were labelled A and B in the experiment. The information that each player received at the end of every round was the visualization of the action chosen in that round (A or B), the payoff gained (1 or 0), the additional information prescribed by the experimental treatment, and his or her cumulative payoff, expressed as a percentage of successful rounds over the total at that point⁴.

The experiment was ran at the Computable and Experimental Economics Laboratory of the University of Trento. Subjects were recruited from different departments of the University of Trento and had never participated in experiments of this type before. The experiment was conducted in six sessions of 20 players each, and it was computerized. Subjects were randomly assigned a seat at the lab computer room upon arrival, and were given written instructions⁵. They could see each other, but no form of communication was allowed during

⁴Subjects were explicitly told that in case all five members chose the same action, they would all earn a payoff of zero, and the computer would display the action not chosen by anyone as the winning action.

⁵The complete set of instructions is available from the authors upon request.

time steps	A	F
25	1,18	1,14
50	1,12	1,04
75	1,00	1,04
100	0,93	0,80

Table 3: Values of σ over time averaged across values of M for the aggregate and full information treatments

time	$M = 1$	$M = 4$	$M = 16$
25	1,18	1,22	1,08
50	1,27	0,98	1
75	1,25	0,89	0,93
100	1,01	0,84	0,75

Table 4: Values of σ over time averaged across information treatments for different values of M

the entire session.

Subjects were told that they would be randomly divided in four groups of five players at the beginning of the experiment and that such division would be known only to the computer program. Then they would participate in a repeated decision-making experiment with the same group of people for a total of 100 periods. Each subject received a fixed show up fee plus anything he or she could earn in the experiment. Payoffs in the game were expressed in ‘experimental points’. At each round a subject could earn either 1 point or zero, and the monetary exchange rate was fixed at E.18 for each point, for a theoretical maximum of E25.82 per subject (show up fee included). When all members of a group entered their choice, the computer calculated each player’s payoff and disclosed information about the round. Sessions lasted fortyfive minutes on average, including instruction time.

4 Results

4.1 Aggregate behavior

We first define as a measure of *allocative efficiency* the squared deviation Σ of the number of agents choosing the 0 side from the half population value:

$$\Sigma = \sum_t (N_0(t) - \frac{N}{2})^2 \quad (2)$$

where $N_0(t)$ is the number of players choosing party 0 at time t and N is the total number of players participating in the game. The value of Σ can be comprised between 0.25 and 6.25. Clearly, the lower its value, the higher the efficiency.

Tab. 2 reports the values of σ over time separately by the value of M and averaged across groups in both the aggregate and full information treatments.

The ‘benchmark’ value for σ in the case of random symmetric players is equal to 1.25. Note that the degree of allocative efficiency in all treatments is significantly higher than the benchmark value. Furthermore, a general increasing trend in efficiency is clearly detectable (see also Tab. 3 and 4).

Hence, all groups in the experiment are able to achieve a quite remarkable degree of self-organization, with a higher efficiency than the one attainable through purely random play. Recall that the highest degree of efficiency would be obtained in correspondence of any of the pure strategy Nash equilibria, whereby players repeatedly choose the same side. However, as such equilibria are not strict and payoff-asymmetric, they were in fact never observed.

Given the high degree of allocative efficiency obtained by all groups, and hence the high level of dynamic coordination between players, one might ask whether such coordination produces ‘predictable’ patterns in the series of aggregate outcomes. In fact, such a good level

	Aggregate Info.			Full Info.		
	Mean	St. dev.	K.S. test	Mean	St. dev.	K.S. test
All	32.6	6.67	0.039	33.15	6.41	0.008
$m = 1$	31.7	5.46	0.56	31.9	5.22	0.56
$m = 4$	33.7	9.13	0.016	32.75	4.89	0.11
$m = 16$	32.6	4.77	0.16	34.8	8.48	0.03
Theoretical	31.25	4.63				

Table 5: Payoff distribution analysis

of coordination could be produced by the players' use of repeated choice patterns that could reflect themselves into regularities at the aggregate level.

Indeed, if this was the case, the observed level of allocative efficiency would not be 'robust' because it could be in principle decreased by some hypothetical arbitraging behaviors aimed at exploiting the predictability in the time series of outcomes.

In order to investigate the issue, we introduce a measure of *informational efficiency*, which describes the amount of 'informational content' of the history of interaction. In fact, if players behave randomly or if there is a significant degree of heterogeneity and change in their behavior, the aggregate information represented by the history of interaction - i.e., the string of successive winning sides at each round - should be perfectly described by a series of i.i.d. draws and hence render the outcome in the next period of play perfectly unpredictable. The resulting 'market' would be, in a sense, perfectly efficient. Otherwise, the presence of 'structure' in the time series would indicate the presence of some 'slack' in the players' coordination patterns that could in principle be exploited.

Note that the degree of informational efficiency of the resulting 'market' is an aggregate measure of closeness to the mixed strategy equilibrium solution. In fact, mixed strategy equilibrium behavior implies that the series of winning sides be a sequence of unrelated draws.

We introduce a measure of *informational efficiency* Λ defined as

$$\Lambda(l) = \sum_{h \in I_l} \max\{F(0|h), F(1|h)\} \quad (3)$$

where I_l is the set of all the binary strings of length l and $N(i|h)$ with $i \in \{0, 1\}$ is the number of times i has been the winning side when h was the preceding history of length l (i.e. the string of the previous l winning sides). From its definition, $\Lambda(l)$ represents the average payoff of the (ex-post defined) best strategy of a hypothetical player who has access to the last l rounds of the history of play. It follows that $\Lambda(l)$ is a non decreasing function of l .

The results for the informational efficiency are shown in Fig. 1 and Fig. 2 for the aggregate and full information treatments respectively. We calculated the values of Λ for different values of the history length l . The dotted lines represent the 1 st.dev. band relative to the usual 'benchmark' case. Interestingly, as the figures show, irrespectively of the information and memory amounts, the observed informational efficiency departs from the 'benchmark' values for less than one standard deviation, suggesting a substantial lack of structure in the outcome time series.

4.2 Analysis of payoffs distribution

The last aggregate analysis concerns the total payoff distribution over the players population in all treatments. The natural benchmark to which compare the observed distribution is again the “random players” case where all the players involved choose randomly with equal probabilities between the two sides.

If N is the number of players the winning side can be of size $0, \dots, (N - 1)/2$ and the probability for any given player of belonging to a winning side of size k is k/n . Then the probability to be a winner can be computed as

$$p_w = \sum_{h=0}^{(N-1)/2} 2 \binom{N}{h} \frac{h}{N} \quad (4)$$

where the factor 2 accounts for the symmetry of the game and the probability of winning k matches out of M is given by a binomial distribution:

$$p(k; M, p_w) = \binom{M}{k} p_w^k (1 - p_w)^{M-k} \quad (5)$$

In our case $N = 5$ and since we are concerned with the total payoff in each treatment, $M = 100$. Tab. 5 shows the mean and average payoff for all treatments and for the subpopulation obtained considering only treatments with a given m . The distributions so obtained are then compared with the theoretical distribution defined in Eq. 5 by using the Kolmogorov Smirnov statistics (von Mises 1964). As can be seen, the null hypothesis of identity between the observed distributions and the theoretical prediction based on purely random players can never be disproved with high significance, even if the variance of the observed payoffs seems larger especially in the $m = 4$ case.

Hence, concluding the section on aggregate behavior, we can summarize the main results as follows: our groups of players were able to achieve a high degree of coordination - and hence were fully able of exploiting the game’s gain opportunities - without producing structure in the times series of aggregate outcomes, i.e., producing a correspondingly high level of informational efficiency. Finally, the analysis of payoffs reveals that their distribution upon the population does not depart significantly from the theoretical benchmark of perfect randomizing players. In the next section we will analyze how players did implement individually the observed collective self-organization.

4.3 Individual Behavior

In this section we analyze individual behavior in the game. First, we check whether choices are biased toward one of the two possible options A and B due to their labels. One may in fact hypothesize that players react to choice labels in various “irrational” ways or in the effort of using them as a coordination device. A binomial test performed on choices pooled across treatments failed to reject the null hypothesis of no association between action label and choice⁶. Hence, the results of the subsequent analysis is collapsed across the label variable.

We next compare individual data with the theoretical benchmark represented by the symmetric mixed strategy equilibrium solution. There is considerable evidence that individuals

⁶Choices of A are 50.2% of all choices (binomial test, $p = .05$). From now on, all statements on statistical significance based on nonparametric tests will always be based on the conventional significance level of .05.

have difficulties in generating random sequences, even when explicitly asked to do so. Many of them, in fact, exhibit a negative recency bias, generating series with too many alternations compared to those of a purely random series (e.g., Budescu 1987). This phenomenon seems related to the representative heuristic (Kahneman, Slovic and Tversky 1982) according to which a binary series with many alternations is generally judged as very representative of a random series, although it is not random. However, in competitive game-theoretic settings in which subjects repeatedly play games with a unique, mixed strategy equilibrium, the performance is generally better, as subjects have an incentive to render their future actions unpredictable (Rapoport and Budescu 1992), although the negative recency bias does not disappear completely (Rapoport and Boebel 1992).

In the minority game, the amount of information available to players might influence the extent to which they behave randomly; in particular, in the full information treatment, in which everybody's choice is disclosed to all group members, players may be more induced to render their choice unpredictable and may try to exploit predictable patterns in others' choices, and this is more likely to occur the higher the value of M (the longer the string of the game history).

However, a longer 'memory' of one's own past performance associated to each choice may induce players to take more into account their personal history of gains and losses when choosing, which might go against randomizing behavior (which intuitively would require a 'memoryless' selection procedure).

In the aggregate information treatment, a binomial test performed on each individual sequence of choices allows to reject the null hypothesis of a binomial distribution in the 35% of cases (5 of 20 players in both the $m = 1$ and $m = 16$ treatments, and 6 of the 20 players in the $m = 4$ treatment). Few of these players followed almost pure strategies, i.e., they almost always chose the same side. The vast majority of the players did not follow pure strategies. However, only few of the mixed strategies can be accounted for by the mixed-strategy equilibrium.

In the full information treatment, the percentage of players' strings for which the binomial distribution can be rejected rise up to 52% (9 of 20 players in the $M=1$ treatment, 5 of 20 players in the $M=4$ treatment, and 17 of 20 players in the $M=16$ treatment), a datum that seem to go against the intuition that increased information would increase randomizing behavior.

Let us then introduce a finer-grained investigation of individual choice patterns. We will next detect the presence of 'structure' in the time series of players' choices, where by 'structure' we mean any deviation from the null hypothesis of mixed strategy equilibrium.

In order to provide a synthetical description of possible general relations between a player's choice and the information he possesses about the game (i.e., his previous choices and the information about past aggregate outcomes and, in the full information treatment, the individual choices of others) we use the measure of *entropy*, which defines the informational content of a probability distribution.

In order to clarify the analysis that follows, let us first review a few formal definitions:

Given a random variable X taking discrete values with probabilities $\{p_1, \dots, p_N\}$ its entropy is defined as

$$H(X) = - \sum_{i=1}^N p_i \log(p_i) \quad (6)$$

and characterizes its information content (see Shannon 1948). In the same way, considering two random variables X and Y , one can compute the entropy of the the joint distribution $H(X, Y)$

l	m=1	m=4	m=16	bench.
1	-4.39 (-6.44,-2.95)	-6.41 (-7.39,-4.53)	-4.74 (-5.24,-3.23)	-6.06 (-7.22,-4.7)
2	-3.22 (-4.19,-2.53)	-4.00 (-4.99,-3.18)	-3.42 (-4.02, -2.38)	-4.14 (-4.68,-3.44)
3	-2.49 (-3.04,-1.89)	-2.84 (-3.33,-2.41)	-2.52 (-3.16,-1.95)	-3.05 (-3.4,-2.64)
4	-1.77 (-2.14,-1.50)	-1.97 (-2.34,-1.69)	-1.74 (-2.08,-1.52)	-2.09 (-2.33,-1.83)
1	-4.94 (-5.45,-3.20)	-4.61 (-5.53,-2.65)	-4.84 (-6.01,-3.60)	-6.06 (-7.22,-4.7)
2	-3.56 (-4.40,-2.99)	-3.41 (-4.11,-2.44)	-3.51 (-4.42,-2.83)	-4.14 (-4.68,-3.44)
3	-2.71 (-3.16,-1.97)	-2.43 (-3.02,-1.94)	-2.34 (-2.78,-1.88)	-3.05 (-3.4,-2.64)
4	-1.94 (-2.31,-1.56)	-1.78 (-2.19,-1.48)	-1.65 (-1.92,-1.50)	-2.09 (-2.33,-1.83)

Table 6: Mean and st. dev. for the informational content U of the string of choices with respect to the string of past outcomes for the aggregate info. (top) and the full info. (bottom) cases. The values for four different depths $l = 1, 2, 3, 4$ are reported.

and the entropy of one variable with respect to the other, for instance of the variable Y with respect to the variable X which is defined as $H(Y|X) = H(X, Y) - H(X)$. This quantity can be thought to represent the increase in information obtained observing the realizations of the variable Y when the variable X is already known. If $H(Y|X) = H(Y)$ the increment in information is exactly the information content of the variable itself, in other words, the previous knowledge of X doesn't reduce the uncertainty on Y and hence the two variables can be considered to be unrelated. Indeed, it is easy to show that the relation $H(X, Y) = H(X) + H(Y)$ holds if and only if X and Y are independently distributed random variables.

Following from the previous properties the ‘‘uncertainty coefficient’’ of a variable Y with respect to a variable X can be defined as:

$$U(Y|X) = \frac{H(Y) - H(Y|X)}{H(Y)} = \frac{H(X) + H(Y) - H(X, Y)}{H(Y)} \quad (7)$$

This ‘‘coefficient’’ represents the fractional contribution to the knowledge of Y given by the previous knowledge of X , i.e. which fraction of the possible realizations of the variable Y can be predicted knowing the actual realizations of X . If $U(Y|X) = 0$ the variables are independent, while if $U(Y|X) = 1$ they are identically distributed; therefore U can be considered a measure of probabilistic dependence between X and Y . Notice that in this method no ad-hoc assumption is made on the nature of the probability distributions and on the form of the relation between them.

For what concerns our analysis, we fix a ‘‘memory depth’’ l and consider the joint probability distribution $p(c, h)$ where $c \in \{0, 1\}$ is a player's choice at a given round and $h \in \times^l \{0, 1\}$ is the past history of winning sides of length l at that time. This distribution can be computed for each player in each treatment from the observed frequencies and analogously can be computed the probability distribution of choices $p(c)$ and of past histories $p(h)$.

Using Eq. (7) we can define an ‘‘uncertainty coefficient’’ $U_{\text{hist}}(l) = U(c|h)$ of choices with respect to previous outcomes for each player in each treatment and obtain, for a given memory m , a distribution of $5 \times 4 = 20$ points. In Tab.6 we compare their means with a ‘‘benchmark’’ distribution obtained by a Monte Carlo simulation using group of 5 perfectly random player. The same procedure is repeated using, instead of the history of past aggregate outcomes, the string of the l past choices of the player himself, obtaining a distribution for $U_{\text{self}}(l)$ compared with the Monte Carlo distribution in Tab.7. Notice that, due to the relatively short length of

l	m=1	m=4	m=16	bench.
1	-3.14 (-3.49,-2.27)	-3.96 (-4.80,-3.21)	-4.36 (-5.12,-2.73)	-6.06 (-7.22,-4.7)
2	-2.44 (-2.99,-2.06)	-2.81 (-3.55,-2.18)	-3.00 (-3.84,-2.40)	-4.14 (-4.68,-3.44)
3	-2.11 (-2.57,-1.80)	-2.13 (-2.76,-1.70)	-2.36 (-2.69,-1.93)	-3.05 (-3.4,-2.64)
4	-1.54 (-1.71,-1.48)	-1.55 (-1.81,-1.42)	-1.75 (-2.02,-1.40)	-2.09 (-2.33,-1.83)
1	-4.94 (-5.45,-3.20)	-4.61 (-5.53,-2.65)	-4.84 (-6.01,-3.60)	-6.06 (-7.22,-4.7)
2	-3.56 (-4.40,-2.99)	-3.41 (-4.11,-2.44)	-3.51 (-4.42,-2.83)	-4.14 (-4.68,-3.44)
3	-2.71 (-3.16,-1.97)	-2.43 (-3.02,-1.94)	-2.34 (-2.78,-1.88)	-3.05 (-3.4,-2.64)
4	-1.94 (-2.31 -1.56)	-1.78 (-2.19,-1.48)	-1.65 (-1.92,-1.50)	-2.09 (-2.33,-1.83)

Table 7: Mean and st. dev. for the informational content U of the string of choices with respect to the string of past choices for the aggregate info. (top) and the full info. (bottom) cases. The values for four different depths $l = 1, 2, 3, 4$ are reported.

the time series, the length parameter l cannot take too large values. Therefore we limit our analysis to $l \leq 4$.

The first thing to notice in Tab. 6 and Tab. 7 is that, irrespectively of l and m , the relationships between a player’s present choice and his own past choices results stronger than the one between present choices and previous outcomes. It might appear that players tend to repeat, quite irrationally, their previous choices rather than trying to spot and exploit possible patterns in the series of group outcomes, a behavior that would be suggested by any “sophisticated” approach to the game. Moreover, no substantial difference can be noticed between treatments, with one exception: the $m = 4$ aggregate information treatment, which is the one characterized by the highest allocative efficiency, is characterized by a substantial lack of “short range” ($l = 1, 2$) relationship between players’ present and past outcomes (see Tab. 6).

A qualitative analysis of choices over time confirms these observations. No statistical differences were found at this level of analysis between treatments⁷. Players in the game choose an action equal to their previous period action, on average, 65% of times, and such percentage is significantly different from the mixed strategy equilibrium hypothesis (binomial test). Hence, players exhibit a certain degree of inertia in their behavior. Such persistence is reinforced when the previous period action has been successful (i.e., when players chose the “winning side”). In fact, in this case players repeat their last period action 73% of the times on average. The percentage after a “losing” round decreases to 61%, but it is still significantly different from the mixed strategy equilibrium hypothesis.

Moreover, the ‘inertia’ and ‘reinforcement’ effects seem to appear even strongly in the full information treatments, as Tab. 9 shows, and they reach their maximal values in correspondence of the $F(M = 16)$ treatment, i.e., in the treatment in which the highest amount of information was given to subjects, as shown in Tab.8. Hence, increasing information seem to increase the tendency toward ‘inertial’ behavior.

On the other hand, players do not seem to react to the previous period “aggregate information”, i.e., to what action was actually the winning one. In fact, players choosing an action equal to the last round winning action is overall equal to 50.3%, which is perfectly consistent with the hypothesis of randomizing behavior.

⁷Kruskal-Wallis and Kolmogorov-Smirnov tests. Each group value was treated as a single independent observation.

	M=1		M=4		M=16	
	A	F	A	F	A	F
$c(t+1)=c(t)$.63 (.11)	.68 (.12)	.62 (.14)	.65 (.14)	.61 (.12)	.74 (.13)
$c(t+1)=c(t)$ if $\pi(t) = 1$.70 (.18)	.77 (.19)	.70 (.18)	.70 (.20)	.70 (.21)	.83 (.14)
$c(t+1)=c(t)$ if $\pi(t) = 0$.59 (.14)	.65 (.13)	.58 (.13)	.61 (.15)	.55 (.11)	.70 (.15)
$c(t+1)=\text{winner}(t)$.49 (.12)	.49 (.08)	.51 (.06)	.49 (.12)	.52 (.09)	.51 (.07)

Table 8: Frequencies of choices of actions that were equal to the previous period action and to the previous period winning action separately for each treatment (standard deviations are reported in parenthesis)

$P\{c(t+1) = c(t)\}$	if $\pi(t) = 1$		if $\pi(t) = 0$	
	A	F	A	F
mean	.70	.76	.57	.65
st.dev.	.13	.18	.07	.15

Table 9: Mean and standard deviation for the probability of choosing an action equal to the previous period action on the pooled population in the aggregate and full information treatments

However, the presence of heterogeneity in players' behavior is likewise revealed, both by the variances reported in Tab. 8 and 9 and, finally, by the analysis reported in Fig. 3, where $p_i(a|b, c)$, the probability that player i chooses $a \in \{0, 1\}$ given his choice was $b \in \{0, 1\}$ in the previous round and the winning side was $c \in \{0, 1\}$, computed starting from the choices string of each player, is plotted. The two tendencies of 'inertia' and 'reinforcement' clearly appear from the plots; however, despite these central tendencies, heterogeneity is relatively high, especially in the (b) and (c) cases.

5 Conclusion and outlooks

The main research question of this work was to investigate, first, the extent to which groups of human players are able to coordinate efficiently in a minority game, and, second, if and to what extent the amount of information available to them makes a difference in terms of both their individual choice patterns and the resulting aggregate efficiency. Our main results can be summarized as follows: at the aggregate level, the degree of both allocative and informational efficiency achieved by our groups are remarkably high in all treatments when compared to the theoretical benchmark represented by the case of symmetric players who follow the unique mixed strategy equilibrium. The lack of a significant difference between treatments suggests that players only need minimal information to coordinate efficiently, and disclosing more information does not necessarily lead to improved aggregate performance.

Second, analysis of individual behavior in the game reveals that such efficiency is not achieved by means of 'rational' (i.e., randomizing) behaviors, but rather through a co-evolution in players' - 'suboptimal' - choice rules. In fact, analysis of individual strings of choices reveals the presence of structure; in particular, players on average exhibit 'inertial' behavior, i.e., the tendency to replicate their previous round choice more often than what would be implied by pure randomization, and this tendency increases when the previous round choice

was successful, suggesting the presence of a ‘reinforcement’ effect (e.g., Erev and Roth 1998).

On the contrary, there is no significant relation between players’ choices and the previous rounds’ ‘winning’ sides, suggesting that players focus relatively more on their personal series of choices than on the behavior of others. Evidence on reinforcement effects confirms previous experimental studies on market entry games showing that subjects seemed to be affected in their current choices by their personal histories of gains and losses (Zwick and Rapoport 1999).

An additional treatment whereby groups’ compositions vary from round to round (random assignment) would allow to measure the extent to which the high level of coordination observed rests upon some forms of “self-organization”, generated by players’ ability to learn over time to adapt to one another.

In addition, further research might involve a sound confrontation with experimental data from similar coordination games (like the aforementioned market entry games), tests of learning models and further simulations studies in order to pin down the behavioral and institutional conditions that favor the emergence of spontaneous, decentralized coordination in this large class of interactive settings.

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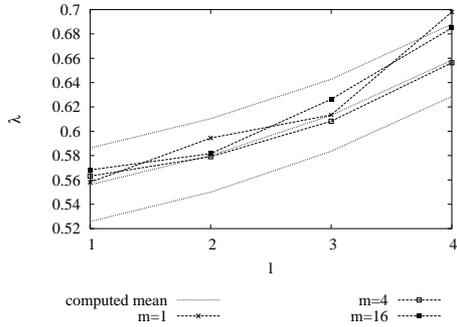


Figure 1: Average informational efficiency for $m = 1, 4, 16$ computed with different values of l in the aggregate information case.

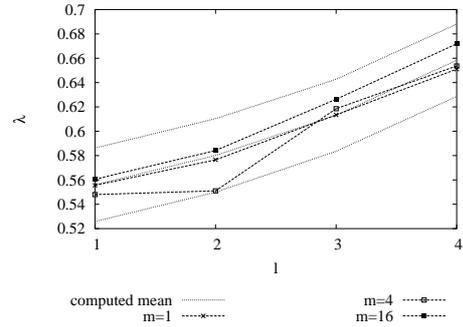


Figure 2: Average informational efficiency for $m = 1, 4, 16$ computed with different values of l in the full information case.

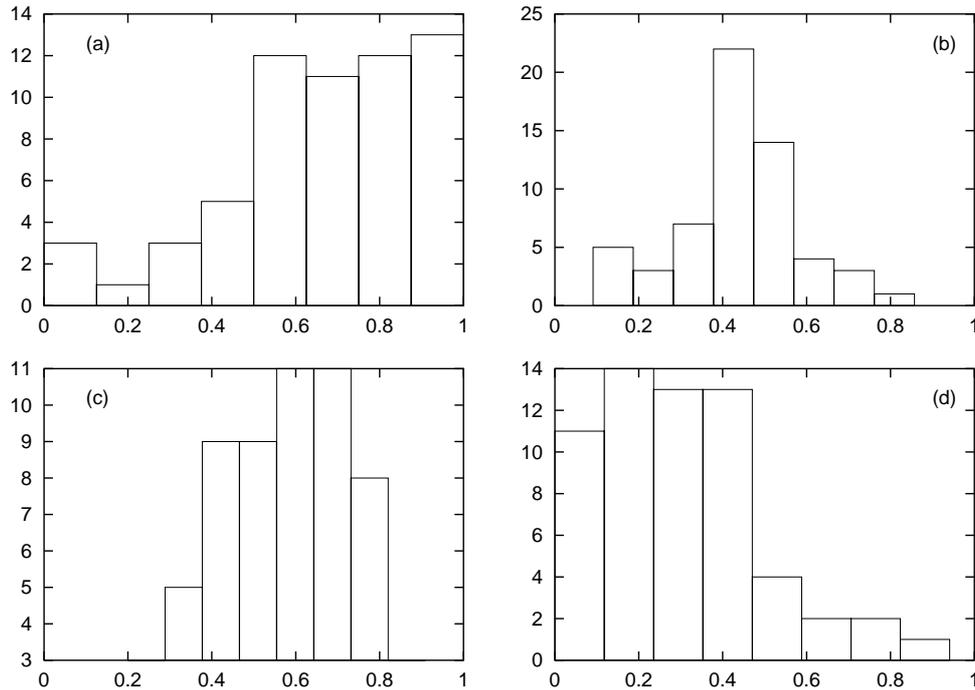


Figure 3: Binned frequency for $p(0|a; b)$. **(a)** When the last choice was 0 and the winning side 0. **(b)** When the last choice was 1 and the winning side 0. **(c)** When the last choice was 0 and the winning side 1. **(d)** When the last choice was 1 and the winning side 1. The persistence is made stronger by a “reinforcement” effect